## Nowcasting Macroeconomic Variables with a Sparse Mixed Frequency Dynamic Factor Model –Supplementary Material–

Domenic Franjic<sup>∗</sup> Karsten Schweikert†

University of Hohenheim

University of Hohenheim

[Latest update: October 30, 2024]

## **A Additional Simulation Results**



Note: *ρ* refers to the average absolute correlation-coefficients of the measurement errors. MSNE reductions are computed as  $1-S$ , where S is given by  $S = \sum_{i=1}^{500} (\hat{x}_{T,0,i,\text{SDFM}} - x_{T,0,i})^2 / \sum_{i=1}^{500} (\hat{x}_{T,0,i,\text{DFM}} - x_{T,0,i})^2$ , where  $\hat{x}_{T,0,i,\text{SDFM}}$  represents the *i*th nowcast of the SDFM model,  $\hat{x}_{T,0,i,\text{DFM}}$  represents the *i*th nowcast of the benchmark DFM model, and  $x_{T,0,i}$ represents the realisation of the variable of interest at time point *T*.

Table 1: MSNE reduction over 1000 simulated nowcasting exercises restricted to a single penalty for all factors using  $T = 100$  observations for  $N = 50$  variables

<sup>∗</sup>Address: University of Hohenheim, Core Facility Hohenheim & Institute of Economics, Schloss Hohenheim 1 C, 70593 Stuttgart, Germany, e-mail: *franjic@uni-hohenheim.de*

<sup>†</sup>Address: University of Hohenheim, Core Facility Hohenheim & Institute of Economics, Schloss Hohenheim 1 C, 70593 Stuttgart, Germany, e-mail: *karsten.schweikert@uni-hohenheim.de*

			$s = 0.00$		$s = 0.80$			
		$\bar{\rho} \approx 0.2 \quad \bar{\rho} \approx 0.3 \quad \bar{\rho} \approx 0.4 \quad \bar{\rho} \approx 0.5 \quad \bar{\rho} \approx 0.2 \quad \bar{\rho} \approx 0.3 \quad \bar{\rho} \approx 0.4 \quad \bar{\rho} \approx 0.5$						
		$R = 1$ 0.411 0.415 0.481		$0.486$   $0.513$		$0.491$ $0.473$		0.450
$R=2$	0.403		$0.404$ $0.516$		$0.509$ 0.490	0.497	0.527	0.505
$R=3$	0.375	0.436	0.480	0.501	0.412	0.420	0.484	0.523
$R=4$	0.433		$0.452$ $0.479$	$0.515$	0.447	0.470	0.491	0.522

Note: The SDFM model hyper-parameters are validated optimising the BIC. *ρ* refers to the average absolute correlation-coefficients of the measurement errors. The ratios of relative MSNE reductions are computed as  $\frac{1}{500} \sum_{i=1}^{500} \mathbb{1} \left( (\hat{x}_{T,0,i,\text{SDFM}} - x_{T,0,i})^2 < (\hat{x}_{T,0,i,\text{DFM}} - x_{T,0,i})^2 \right)$ , where  $\hat{x}_{T,0,i,\text{SDFM}}$  represents the *i*th nowcast of the SDFM model,  $\hat{x}_{T,0,i,\text{DFM}}$  represents the *i*th nowcast of the benchmark DFM model, and  $x_{T,0,i}$  represents the realisation of the variable of interest at time point *T*.

Table 2: Ratio of relative MSNE reductions over 1000 simulated nowcasting exercises restricted to a single penalty for all factors using  $T = 200$  observations for  $N = 100$  variables

			$s = 0.00$		$s = 0.80$			
					$\bar{\rho} \approx 0.2 \quad \bar{\rho} \approx 0.3 \quad \bar{\rho} \approx 0.4 \quad \bar{\rho} \approx 0.5 \mid \bar{\rho} \approx 0.2 \quad \bar{\rho} \approx 0.3 \quad \bar{\rho} \approx 0.4 \quad \bar{\rho} \approx 0.5 \mid$			
$R=1$		$-0.483 -0.225 -0.212$		-0.034	0.017	-0.194	$-0.008$	$-0.022$
$R=2$	$-0.424$	$-0.049$	$-0.005$	$-0.001$	$-0.024$	$-0.081$	$-0.050$	$-0.109$
$R=3$	$-0.413$	0.010	$-0.147$	0.012	$-0.493$	$-0.089$	0.034	0.024
$R=4$	$-0.358$	0.022	0.014	0.013	$-0.313$	0.038	0.037	$-0.060$

Note: The SDFM model hyper-parameters are validated optimising the BIC. The SDFM model hyper-parameters are validated optimising the BIC.  $\rho$  refers to the average absolute correlation-coefficients of the measurement errors. MSNE reductions are computed as  $1-S$ , where S is given by  $S = \sum_{i=1}^{500} (\hat{x}_{T,0,i,\text{SDFM}} - x_{T,0,i})^2 / \sum_{i=1}^{500} (\hat{x}_{T,0,i,\text{DFM}} - x_{T,0,i})^2$  $(x_{T,0,i})^2$ , where  $\hat{x}_{T,0,i,\text{SDFM}}$  represents the *i*th nowcast of the SDFM model,  $\hat{x}_{T,0,i,\text{DFM}}$  represents the *i*th nowcast of the benchmark DFM model, and  $x_{T,0,i}$  represents the realisation of the variable of interest at time point *T*.

Table 3: MSNE reduction over 1000 simulated nowcasting exercises restricted to a single penalty for all factors using  $T = 200$  observations for  $N = 100$  variables



Note: The SDFM model hyper-parameters are validated optimising the BIC. *ρ* refers to the average absolute correlation-coefficients of the measurement errors. The ratios of relative MSNE reductions are computed as  $\frac{1}{500} \sum_{i=1}^{500} \mathbb{1} ((\hat{x}_{T,0,i,\text{SDFM}} - x_{T,0,i})^2 < (\hat{x}_{T,0,i,\text{DFM}} - x_{T,0,i})^2)$ , where  $\hat{x}_{T,0,i,\text{SDFM}}$  repr SDFM model,  $\hat{x}_{T,0,i,\text{DFM}}$  represents the *i*th nowcast of the benchmark DFM model, and  $x_{T,0,i}$  represents the realisation of the variable of interest at time point *T*.

Table 4: Ratio of relative MSNE reductions over 500 simulated nowcasting exercises restricted to a single penalty for all factors using  $T = 200$  observations for  $N = 100$  Variables

## **B Algorithms**

In the upcoming section, the following conventions are used. The symbol ⊔ denotes the operation of adding an element to the end of a sequence. Let  $\mathcal S$  be a sequence of integer values and  $\mathbf{A} = (a_{n,m})_{n=1,m=1}^{N,M}$  be a matrix. Assuming min  $\mathcal{S} \geq 1$  and max  $\mathcal{S} \leq M$ ,  $\mathbf{A}[\mathcal{S}, \cdot]$ refers to the submatrix of **A** that is constructed by concatenating the rows with index corresponding to the integers in S in order. For a vector  $\mathbf{v} = (v_n)_{n=1}^N$ , and  $\min S \ge 1$  and  $\max S \leq N$ ,  $\mathbf{v}[S]$  denotes the subvector consisting of the elements of **v** with corresponding index  $n \in \mathcal{S}$ . Similarly,  $\mathbf{v}[n_1 : n_m]$  refers to the subvector consisting of the elements with index  $n_1$  to index  $n_m$ . Adding an element v at the end of a vector is denoted as  $[\mathbf{v}, v]$ . Analogously,  $[\mathbf{v}_1, \mathbf{v}_2]$  concatenates two vectors  $\mathbf{v}_1, \mathbf{v}_2$ . Initialising an empty vector is denoted as  $\mathbf{v} = [\cdot]$ . The operation of removing an element  $v_n$  from a vector **v** is defined as  $\mathbf{v}_1, \mathbf{v}_2$  ( $v_1, \ldots, v_{n-1}, v_{n+1}, \ldots, v_N$ ). For two vectors  $\mathbf{v}_1, \mathbf{v}_2$  of equal length, the element-wise division is defined as  $\mathbf{v}_1 \oslash \mathbf{v}_2 = (v_{1,1}/v_{1,2}, \ldots, v_{n,1}/v_{n,2}).$ 

**Algorithm 1** Sparse Principal Components Analysis [\(Zou and Hastie,](#page-5-0) [2020\)](#page-5-0)

1:  $X' = UDV'$ 

2:  $\mathbf{A} \leftarrow (\mathbf{v}_1, \dots, \mathbf{v}_R)$ , where  $\mathbf{v}_1, \dots, \mathbf{v}_R$  correspond to the first *R* columns of **V**.

3: Set **A** to a Matrix with entries equal to the double precision floating point maximum

4: Set conversion threshold  $\epsilon > 0$ 

5: **while**  $\|\mathbf{A} - \mathbf{A}\|_F > \epsilon$  do

$$
6: \qquad \tilde{\mathbf{A}} \leftarrow \mathbf{A}
$$

- 7: **for**  $r \in \{1, ..., R\}$  **do**
- 8: Use Algorithm [2](#page-2-0) to solve

$$
\widehat{\boldsymbol{\lambda}}_r \leftarrow \operatorname*{argmin}_{\boldsymbol{\lambda}_r} \left\{ \left( \boldsymbol{\alpha}_r - \boldsymbol{\lambda}_r \right)' \mathbf{X}_{\tau} \mathbf{X}_{\tau}' \left( \boldsymbol{\alpha} - \boldsymbol{\lambda}_r \right) + \kappa_2 \| \boldsymbol{\lambda}_r \|_2 + \kappa_{1,r} \| \boldsymbol{\lambda}_r \|_1 \right\} \tag{B.1}
$$

- 9: **end for**
- 10:  $\boldsymbol{\lambda} \leftarrow (\boldsymbol{\lambda}_1, \ldots, \boldsymbol{\lambda}_R)$
- 11: Compute the singular value decomposition of  $\mathbf{X}\mathbf{X}'\hat{\lambda}$ , i.e.,  $\mathbf{X}\mathbf{X}'\hat{\lambda} = \mathbf{U}\mathbf{D}\mathbf{V}'$ .
- 12: **A** ← **UV**<sup> $\prime$ </sup>
- 13: **end while**
- 14: **return**  $\widehat{\Lambda} = \widehat{\lambda}$

<span id="page-2-0"></span>**Algorithm 2** Least Angle Regression [\(Efron et al.,](#page-5-1) [2004;](#page-5-1) [Zou and Hastie,](#page-5-2) [2005,](#page-5-2) [2020\)](#page-5-0)

1: Set *N* to the number of rows of **X**

2: Set *T* to the number of columns of **X**

3: 
$$
\mathbf{y} \leftarrow \mathbf{X} \alpha_r
$$
  
4:  $\mathbf{c} \leftarrow \left| \frac{1}{\sqrt{1+\kappa_2}} \mathbf{X} \mathbf{y} \right|$ 

5:  $p \leftarrow \max c$ 6:  $\mathcal{I} \leftarrow (1, \ldots, N)$ 7:  $\mathcal{A} \leftarrow \emptyset$ 8:  $d \leftarrow 0$ 9:  $\mathbf{s} \leftarrow [\cdot]$ 10:  $\boldsymbol{\lambda}_r \leftarrow \mathbf{0}_N$ 11:  $\hat{\mathbf{q}} \leftarrow \mathbf{0}_N$ 12: Set *E* to double precision floating point maximum 13: Set threshold values *M* for  $0 < M \leq N$  and/or  $\kappa_{1,r}$ 14: **while** 2*p* √  $\overline{1+\kappa_2} > \kappa_{1,r} \wedge |\mathcal{A}| < M$  do 15:  $\hat{c} \leftarrow \max |\mathbf{x}_n \mathbf{y}| \text{ for } n \in \mathcal{I} \text{ and } \mathbf{x}_n \in \mathbf{X}$ <br>
16:  $\hat{n} \leftarrow \operatorname{argmax} |\mathbf{x}_n \mathbf{y}| \text{ for } n \in \mathcal{I} \text{ and } \mathbf{x}_n \in \mathbf{X}$ 16:  $\hat{n} \leftarrow \operatorname{argmax} |\mathbf{x}_n \mathbf{y}| \text{ for } n \in \mathcal{I} \text{ and } \mathbf{x}_n \in \mathbf{X}$ <br>17: **if**  $d = 0 \land |\mathcal{A}| < N$  **then** if  $d = 0 \land |\mathcal{A}| < N$  then 18:  $\mathcal{A} \leftarrow \mathcal{A} \sqcup {\hat{n}}$ <br>19:  $\mathcal{I} \leftarrow \mathcal{I} \setminus {\hat{n}}$  $\mathcal{I} \leftarrow \mathcal{I} \backslash {\hat{n}}$ 20: **s** ←  $[\mathbf{s}, \text{sign}(\widehat{c})]$ <br>21: **if**  $|\mathcal{A}| = 1$  **the**: if  $|A| = 1$  then 22: **L** ←  $\sqrt{\mathbf{x}_n \mathbf{x}_n + \kappa_2}/(1 + \kappa_2)$ 23: **else** 24: Update the lower Cholesky matrix **L** by  $\mathbf{x}_{\hat{n}}$ 25: **end if** 26: **end if** 27: **if**  $d = 1 \vee |\mathcal{A}| = N$  **then** 28:  $d \leftarrow 0$ 29: **end if** 30:  $\mathbf{g} \leftarrow \mathbf{s}$ 31: **g**  $\leftarrow$  **v**<sub>1</sub>, where **v**<sub>1</sub> is the solution to  $Lv_1 = g$ 32: **g**  $\leftarrow$  **v**<sub>2</sub>, where **v**<sub>2</sub> is the solution to  $\mathbf{L}'\mathbf{v}_2 = \mathbf{g}$ 33:  $a \leftarrow 1/$ √ **g** ′**s** 34:  $\mathbf{w} \leftarrow a\mathbf{g}$ 35: **u** ←  $\frac{1}{\sqrt{1}}$  $\frac{1}{1+\kappa_2}\mathbf{X}[\mathcal{A},\cdot]' \mathbf{w}, \frac{\sqrt{1+\kappa_2}}{\sqrt{1+\kappa_2}}$ √ *κ*2  $\frac{\sqrt{\kappa_2}}{1+\kappa_2} \mathbf{w}\Big]$ 36:  $\boldsymbol{\gamma} \leftarrow -1 \cdot (\boldsymbol{\lambda}_r[\mathcal{A}] \oslash \mathbf{w})$ 37: **if**  $0 < \max \gamma$  **then** 38:  $\tilde{\gamma} \leftarrow \max \gamma$ 39:  $n \leftarrow \operatorname{argmin} \gamma_i \text{ for } \gamma_i \in \gamma$ 40: **else** 41:  $\tilde{\gamma} \leftarrow E$ 42: **end if**

43:  $\hat{\gamma} \leftarrow \hat{c}/a$ 44: **if**  $|\mathcal{A}| < N$  **then** 45:  $\alpha \leftarrow \mathbf{X}[\mathcal{A}, \cdot] \mathbf{u}[1:T] + \sqrt{\kappa_2} \mathbf{u}[(T+1):|\mathcal{A}|\right]$ 46: **for**  $m \in (1, \ldots, |\mathcal{I}|)$  **do** 47:  $h_1 \leftarrow (\hat{c} - c_{i_m})/(a - \alpha_m)$ , where  $i_m \in \mathcal{I}, c_{i_m} \in \mathbf{c}$ , and  $\alpha_m \in \mathbf{\alpha}$ 48:  $h_2 \leftarrow (\hat{c} + c_{i_m})/(a + \alpha_m)$ , where  $i_m \in \mathcal{I}$ ,  $c_{i_m} \in \mathbf{c}$ , and  $\alpha_m \in \mathbf{\alpha}$ <br>49: **if**  $0 < h_1 < \hat{\gamma}$  **then** 49: **if**  $0 < h_1 < \hat{\gamma}$  **then**<br>50:  $\hat{\gamma} \leftarrow h_1$ 50:  $\hat{\gamma} \leftarrow h_1$ <br>51: **end if** 51: **end if** 52: **if**  $0 < h_2 < \hat{\gamma}$  then<br>53:  $\hat{\gamma} \leftarrow h_2$ 53:  $\widehat{\gamma} \leftarrow h_2$ <br>54: **end if** 54: **end if** 55: **end for** 56: **end if** 57: *γ*  $\gamma^{\dagger} \leftarrow \min\{\tilde{\gamma}, \hat{\gamma}\}$ 58: λ  $^\dagger \leftarrow \widehat{\boldsymbol{\lambda}}_r$ 59:  $r \leftarrow \widehat{\boldsymbol{\lambda}}_r + \gamma^{\dagger}\mathbf{w}$ 60:  $\hat{\mathbf{q}} \leftarrow \hat{\mathbf{q}} - \gamma^{\dagger} \mathbf{u}$ 61: *p*  $p^{\dagger} \leftarrow p - |\gamma a|$ 62: **if**  $2p^{\dagger}\sqrt{1 + \kappa_2} > \kappa_{1,r}$  then 63:  $q_1 \leftarrow 2p^{\dagger} \sqrt{1 + \kappa_2}$ 64:  $q_2 \leftarrow 2p$ √  $\sqrt{1 + \kappa_2}$ 65:  $\hat{\lambda}_r \leftarrow \left(\frac{q_2 - \kappa_{1,r}}{q_2 - q_1}\right)$  $\frac{q_2-\kappa_{1,r}}{q_2-q_1} \widehat{\boldsymbol{\lambda}}_r + \frac{\kappa_{1,r}-q_1}{q_2-q_1}$  $\left( \frac{q_1,r-q_1}{q_2-q_1} \boldsymbol{\lambda}^{\dagger} \right)$  ( √  $\sqrt{1 + \kappa_2}$ <sup>-1</sup> 66:  $p \leftarrow p^{\dagger}$ 67: **else** 68: **if**  $\tilde{\gamma} < \hat{\gamma}$  **then** 69:  $d \leftarrow 1$ 70:  $\lambda$  $\hat{\lambda}_r[n] \leftarrow 0$ 71:  $\mathcal{I} \leftarrow \mathcal{I} \sqcup \{n\}$ 72: **s** ←  $\left| \mathbf{s}, s_{|\mathcal{A}|} \right|$ 73:  $\mathcal{A} \leftarrow \mathcal{A} \setminus \{n\}$ 74: Downdate the lower Cholesky matrix **L** by **x***<sup>n</sup>* 75: **end if** 76: **end if** 77: **end while** 78: **return**  $\lambda_r$ 

## **References**

- <span id="page-5-1"></span>Efron, B., Hastie, T., Johnstone, I., and Tibshirani, R. (2004). Least angle regression. *The Annals of Statistics*, 32(2):407 – 499.
- <span id="page-5-2"></span>Zou, H. and Hastie, T. (2005). Regularization and variable selection via the elastic net. *Journal of the Royal Statistical Society: Series B (Statistical Methodology)*, 67(2):301–320.
- <span id="page-5-0"></span>Zou, H. and Hastie, T. (2020). *elasticnet: Elastic-Net for Sparse Estimation and Sparse PCA*. R package version 1.3.